

Dynamic Dependence in Corporate Credit

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April 2013

- The CDS market has grown enormously over the last decade
 - \$0.6 trillion of gross notional principal in 2001
 - \$25.9 trillion at the end of 2011 (ISDA)
- The authors motivate their analysis of the dependence between CDS spreads via potential applications to:
 - 1 Managing portfolio credit risk
 - 2 Pricing structured credit products
 - 3 Including credit risk products in an investment portfolio

Overview of the paper

- The authors use weekly CDS spread data on 223 North American firms, March 2003–Sep 2012.
 - These are all firms that appeared in the first 18 CDX IG series
 - Their estimation method allows them to easily handle missing and non-overlapping data
- The authors propose a ~~DCC correlation model~~ skew t copula with dynamics governed by a DCC-type structure. This enables them to exploit two neat results from the DCC literature:
 - 1 Correlation targeting: the $N \times N$ “intercept” matrix can be estimated analytically rather than numerically
 - 2 The model can be estimated using “composite maximum likelihood”
 - These make estimation of a 223-dimensional dynamic copula feasible.

From the lit review section of my own paper

- Papers using time-varying and high ($N \geq 10$) dimension copulas:

Authors	N	Copula	Dynamics	Estim
Zhang, <i>et al.</i> (2011, wp)	10	Skew t	GAS	ML
Christoffersen, <i>et al.</i> (2012, RFS)	33	Skew t	DCC	ML
Almeida, <i>et al.</i> (2012, wp)	30	Vine	SV	SML
Stöber and Czado (2012, wp)	10	Vine	RS	Bayes
Oh and Patton (2013, wp)	100	Factor	GAS	ML

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Oh and Patton (2013, wp)	100	Factor	GAS	ML
Christoffersen, et al. (2013, wp)	233	Skew t	DCC	CML

Main empirical findings of this paper

- 1 The conditional dependence (copula) between CDS spreads varies substantially through time
 - In addition to week-to-week variation, the authors find that overall dependence rose after the financial crisis
 - This reduces potential diversification benefits across these contracts
- 2 The conditional copula of CDS spreads is non-Normal
 - Strong evidence of non-zero tail dependence \Rightarrow not Normal copula
 - Milder evidence of asymmetric dependence, and less evidence for CDS spreads than for equities

1 Motivation and main contributions of paper

2 Suggestions for the authors

- 1 The CDS Big Bang
- 2 Cross-contract spill-overs
- 3 CDS spreads vs. implied probabilities

Outline of this discussion

1 Motivation and main contributions of paper

2 **Suggestions for the authors**

1 The CDS Big Bang

2 Cross-contract spill-overs

3 CDS spreads vs. implied probabilities

The CDS “Big Bang” I

- With the growth of the CDS market through the 2000s, participants wanted more homogeneous contracts to increase liquidity
- On April 8, 2009, the North American CDS market underwent changes to contract conventions
 - CDS coupons were fixed to be 100 or 500 bp, with upfront payments adjusted accordingly
 - More rigid rules on triggers for “credit events” and auctions that follow such events
 - Move towards central clearing, away from OTC trading
- These changes could potentially change the dynamics of CDS spreads

The CDS “Big Bang” II

- In my paper using similar data (100 firms, **daily** spreads over 2006-2012) we found breaks in the conditional mean and variance for 39 and 66 firms respectively
- I tested for a structural break in an AR(1)-GARCH(1,1) model for these 100 firms using **weekly** data, to match this paper and found:
 - Breaks in the conditional mean for 11 firms
 - Breaks in the conditional variance for 55 firms
- ★ The authors have a larger cross-section and use slightly different mean and variance models, but this is suggestive that allowing for a break may be important

- An important aspect of applying Sklar's theorem to conditional distributions is making sure the information sets used for the margins and conditional copula are the same:

$$\mathbf{F}(y_1, \dots, y_N | \mathcal{F}) = \mathbf{C}(F_1(y_1 | \mathcal{F}), \dots, F_N(y_N | \mathcal{F}) | \mathcal{F})$$

- If we use \mathcal{F}_1 for the 1st margin, \mathcal{F}_N for the N^{th} margin, and \mathcal{F}_c for the copula, then the resulting function $\mathbf{F}(\cdot | \cdot)$ will *not*, in general, be a conditional joint distribution.
 - The authors use ARMA(2,2) - NGARCH models for each marginal distribution, which is potentially a problem. But:

Cross-contract spill-overs II

- If we define $\mathcal{F} = \sigma(\mathcal{F}_1, \dots, \mathcal{F}_N, \mathcal{F}_c)$, **and** it is true that

$$\begin{aligned} Y_1 | \mathcal{F}_1 &\stackrel{d}{=} Y_1 | \mathcal{F} \\ &\vdots \\ Y_N | \mathcal{F}_N &\stackrel{d}{=} Y_N | \mathcal{F} \\ (U_1, \dots, U_N) | \mathcal{F}_c &\stackrel{d}{=} (U_1, \dots, U_N) | \mathcal{F} \end{aligned}$$

then we can effectively use the information sets $(\mathcal{F}_1, \dots, \mathcal{F}_N, \mathcal{F}_c)$ but interpret the results as having used \mathcal{F} throughout, satisfying the condition on the information set.

- Practically, this means we should *test* whether including information from “other” variables is needed. If not, we may exclude it.

Cross-contract spill-overs III

- Of course, using a 233-dimensional VAR is not feasible (or sensible).
- One practical alternative is to include a lagged “market” return to capture some cross-asset spillover.
- In Oh and Patton’s study of **daily** CDS spreads we found:
 - 98 / 100 firms had significant spill-over in the mean
 - 13 / 100 firms had significant spill-over in the variance
- I replicated this on **weekly** data (to match the authors) using an AR(1)-GARCH(1,1) and found:
 - 6 / 100 firms had significant spill-over in the mean
 - 7 / 100 firms had significant spill-over in the variance

★ So I think the authors are OK here

CDS spreads vs. implied default probabilities I

- The authors do a very nice job contrasting the properties of the dependence structure of their 233 CDS spreads log-differences with that of the equity returns of the corresponding firms
 - eg: less asymmetric dependence in CDS spreads than equity returns, but a greater change in the level of dependence around the crisis
- The authors also contrast the dependence structure of CDS spreads log-differences with those of *default intensities*, i.e., CDS-implied probabilities of default.
 - They find very similar results. This is not so surprising:

CDS spreads vs. implied default probabilities II

- Under some simplifying assumptions (see Carr and Wu, RFS, for eg) it is possible to show that a CDS spread (S_t) is given by:

$$S_t = P_t \times LGD_t$$

where P_t is the implied probability of default and LGD_t is the loss-given-default.

- This simple expression can also be obtained as a first-order approximation of more complicated formulas when $P_t \approx 0$.
- The LGD is often taken as fixed (at 0.60) and so:

$$\Delta \log S_t = \Delta \log P_t$$

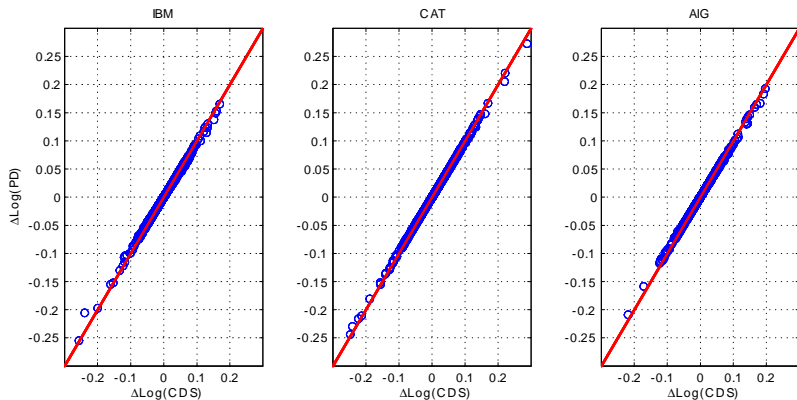
- Thus these two series will be identical.

- What if we use a more serious pricing formula?
 - Conrad, Dittmar and Hameed (2011, wp):

$$S_t = \frac{\sum_{j=1}^{20} d_{t+j|t} (1 - P_t)^j P_t LGD_t}{\frac{1}{4} \sum_{j=1}^{20} d_{t+j|t} (1 - P_t) + \frac{1}{8} \sum_{j=1}^{20} d_{t+j|t} (1 - P_t)^j P_t}$$

- It is not clear what the relation between $\Delta \log S_t$ and $\Delta \log P_t$ will be for this formula. Here are three scatter plots:

CDS spreads vs. implied default probabilities IV



★ $\Delta \log S_t \approx \Delta \log P_t \Rightarrow$ study either $\Delta \log S_t$ or $\Delta \log P_t$?

- A very interesting paper:
 - 1 A large and relatively novel data set
 - 2 A new dynamic copula model, the largest to date
 - 3 Interesting contrasts between the dependence between CDS spreads and dependence between corresponding equity returns